

Application of Machine Learning Algorithms to Detect Treatment Effect Heterogeneity for Three-Level Multisite Experiments

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Motivation (1)

- Multilevel randomized controlled trials (MRCTs) have been widely in education
	- Average treatment effect (ATE)
	- Heterogenous treatment effects (HTE)
		- Under what conditions for whom an intervention works
- Moderation analysis is traditionally used to evaluate HTE
	- An interaction between treatment and moderator (e.g., students' or schools' features)
	- Moderators are usually pre-specified which covariates modify ATE?
- Recent development includes the use of machine learning (ML) methods to explore the HTEs
	- Identify subgroups with *significant effects*, post-analysis what is the expected treatment effect for a group/individual with given characteristics
	- Select moderators from a potentially large number of covariates

Motivation (2)

- MRCTs have nested data structures
	- E.g., students nested within teachers nested with schools
	- Cluster design treatment at the school level
	- Block/multisite design treatment at the student or teacher level
- Observations in the same clusters/sites are correlated rather than independent
	- Multilevel models, cluster robust SEs, bootstrap, etc. have been used to address the dependency
- Similarly, when applying ML methods to estimate CATE, applied researchers still need to consider the nested data structure
	- Most prior literature assumes the participants are independent
	- There is a lack of literature to guide educational researchers in appropriately applying ML methods for clustered data when evaluating HTEs

Purpose

- Review the current available ML methods and tools that account for the nested data structure when explore HTEs
	- Focus on two ML methods Cluster-Robust Causal Forest (CF) & Generic ML
- Demonstrate the application of these two methods using the dataset from a large multisite experimental study (Leite et al., 2023)
	- Provide recommendations to applied researchers on how to choose the appropriate methods and statistical package among alternative ML methods

An Illustrative Example

- A large-scale multisite field experiment embedded within a virtual learning environment (VLE) for Algebra
	- Examine the effects of a video recommendation system
	- Students were randomly assigned to receive two types of video recommendations
		- Personalized video recommendations or generic recommendations
	- 2995 students nested within 54 teachers from 42 schools in three large school districts
	- Measures: 216 student- and teacher-level variables
		- Teacher survey usage of VLE, instructional practice, perception of disruptions due to COVID, etc. (full survey available at [https://osf.io/h5tpn/\)](https://osf.io/h5tpn/)
		- Student variables gender, ethnicity, pretest score, absent days, etc.
		- Converted into 516 predictors, with 484 dummy-coded indicators
			- Some algorithm requires dummy-coding categorical variables

Estimands of Interest

• Conditional average treatment effect (CATE)

$$
\tau(x) = E[Y_{ij}(1) - Y_{ij}(0) | X = x] = u_1(x) - u_0(x) \tag{1}
$$

- $X a$ possibly high-dimensional vector of covariates
- Require stable unit treatment values, unconfoundedness, and overlap
- Best Linear Predictor (BLP) of CATE
- Note that, $ATE = \tau = E[\tau(x)]$, site average or individual average; once we know CATE, we immediately know ATE
- Non-overlapping subgroup analysis
	- E.g., Sorted Group Average Treatment Effects (GATES)

$$
\tau(G_k) = E[Y_{ij}(1) - Y_{ij}(0) | G_k] (2)
$$

• Moderator effect

$$
\tau(m) = E[Y_{ij}(1) - Y_{ij}(0) | M = m] \tag{3}
$$

• *M* is a subset of predictors of *X*

Models: OLS with Teacher Fixed Effects and Interactions

• Traditional Moderator Analysis: Interaction approach

 $y_{ij} = \gamma_0 + \gamma_1 T_{ij} + \gamma_2 M_{ij} + \gamma_3 T_{ij} M_{ij} + u_i + T_{ij} u_j + r_{ij}$ (4)

- y_{ij} test score for student i in teach *j*
- T_{ij} treatment indicator
- M_{ij} student-level moderator
- u_i teacher dummy variables
	- MLM teacher-level random effect that follows a normal distribution
- r_{ij} level-1 error
- γ_3 moderator effects, not a causal effect [\(Dong et al., 2022](https://www-tandfonline-com.lp.hscl.ufl.edu/doi/full/10.1080/00273171.2022.2046997))
- $\tau(M_{ii}) = \gamma_1 + \gamma_3 M_{ii}$, a causal effect
- Assume clusters have an additive effect on the outcome
	- Same functions for all sites

Models: S/T-learner

• Fit *(separate)* models to the treatment and control groups

$$
E(Y|T = 1, X = x) = f_1(x)
$$

$$
E(Y|T = 0|X = x) = f_0(x)
$$

• Then, CATE is

$$
\tau(x) = f_1(x) - f_0(x)
$$

$$
Var(\hat{\tau}(x)) = Var(\hat{f}_1(x)) + Var(\hat{f}_0(x))
$$

• To our best knowledge, no methods or tools consider the nested data structure for S/T-learner

Models: Cluster-Robust RF

• Cluster-robust RF (Athey & Wager, 2019)

$$
y_{ij} = \alpha_j(x) + \tau_j(x)T_{ij} + e_{ij}, \tau(x) = E[\tau_j(x)] \tag{5}
$$

- $\alpha_i(x)$ control group average for site (e.g., teacher) j
- $\tau_i(x)$ CATE in site (e.g., teacher) *j*
- $\tau(x)$ CATE across sites; site average
	- Give each cluster/site equal weight
	- Accurate for predicting effects on a new student from a new site
- Each cluster has its own main $(\alpha_i(x))$ and treatment effect function $(\tau_i(x))$

Models: Generic ML

• Generic ML (Chernozhukov et al., 2023)

$$
y_{ij} = b_0(x) + T_{ij} s_0(x) + e_{ij}, \tau(x) = s_0(x)
$$
 (6)

- $b_0(x) = E[y_{ij} | T_{ij} = 0, x]$ baseline conditional average; mean for the control group across sites
- $s_0(x)$ CATE across level-1 units (e.g., students); individual average
- Site dummy variables can be included when estimating $b_0(x)$ and $s_0(x)$
- An application of double/debiased ML
	- Utilize Neyman orthogonal moments and cross-fitting to address regulation bias and overfitting
- Focus on key features of CATE instead of CATE: e.g., BLP of CATE
	- **Sparsity** s_0 can be well-approximated by a function that only depends on a lowdimensional subset of X

Why/How to Consider Nested Data Structure

- In general, ML methods for CATE include three main steps:
	- Splitting the data into training and test sets cluster-based split?
	- Use the training set and ML algorithms to build a prediction model will considering cluster membership improve prediction?
	- Use the test set to estimate CATE/BLP and their standard errors (SEs) should we use cluster-robust SEs or something similar?
- Based on our review of all the currently available methods and packages, only two algorithms – cluster-robust CF and the GenericML consider the nested data structure in at least one step
	- Not consider alternative methods, e.g., Bayesian additive regression trees (BART), Targeted MLE, Meta learners (e.g., S-, T- learners)

How does Cluster-Robust CF Address the Nesting Effects?

- The cluster-robust CF algorithm considers the nested data structure in *all* three steps:
	- (1) for each b= 1, …, B, draw a **subsample of clusters** and then draw a random sample from each cluster as the training data;
	- (2) grow a tree via recursive partitioning on each such subsample of the data;
	- (3) make the out-of-bag predictions: to account for the potential within cluster dependency, **an observation** i **is considered to be out-of-bag if its cluster was not drawn in step (1)**
- Implemented through *grf* R package
	- SE of CATE jackknife SE
	- Report cluster-robust SEs for BLP

How does the GenericML algorithm Address the Nesting Effects? (1)

- The GenericML algorithm (Chernozhukov et al., 2023) estimates the best linear predictor (BLP) of CATE through the following steps:
	- (1) randomly split the data into training and test sets; without consideration of clusters
	- (2) estimates the CATE with any number of selected ML methods (e.g., random forest) using the training data; *can potentially consider clustering effects*
	- (3) use OLS regression to obtain the BLP of the CATE using the test data; include site fixed effects (dummy variables or demean); easy to report cluster- robust SE from OLS estimation;
- Note that
	- Random forest Build B trees which place covariate splits that maximize the squared difference **in subgroups means**
	- Causal forest Greedily places covariate splits that maximize the squared difference in **subgroup treatment effects**

How does the GenericML algorithm Address the Nesting Effects? (2)

- Implemented through the *GenericML* R package
	- Estimate the Sorted group average treatment effects (GATEs): creating five groups of participants using quintiles of the CATE distribution
	- Perform classification analysis (CLAN) to explore the relationships between covariates and the CATE
	- Report cluster-robust SEs for BLP, GATEs, and CLAN
		- OLS estimation easy to deal with clustering

Results: Cluster-Robust RF

Note. We used lasso, elastic net, support vector machine, XGBoost, and random forests (RF). RF is the best learner.

Blue: CATE estimates from cluster-robust CF Yellow: CATE estimates from CF w/o clustering

Results: GenericML (RF)

Results: BLP of CATE from GenericML (RF)

Black: BLP of CATE estimates without teacher fixed effects

Red: BLP of CATE estimates with teacher fixed effects

Discussion and Conclusion

- Clustering effects should be considered when estimating CATE
	- No perfect solution now
	- It seems CF is preferred
- Differences of CATE estimates between cluster-robust CF and Generic ML
	- CATE or BLP of CATE
	- Site average or individual average
- Future directions
	- Alternative methods: BART, R-learners, TMLE
	- Simulations study
	- Cluster design

Questions or Comments?

Thank you! wei.li@coe.ufl.edu

Appendix: ML and CATE

CR-CF

$$
\hat{\tau}_j = \frac{1}{n_j} \sum_{\{i: A_i = j\}} \hat{\Gamma}_i, \quad \hat{\tau} = \frac{1}{J} \sum_{j=1}^J \hat{\tau}_j, \quad \hat{\sigma}^2 = \frac{1}{J(J-1)} \sum_{j=1}^J (\hat{\tau}_j - \hat{\tau})^2,
$$
\n
$$
\hat{\Gamma}_i = \hat{\tau}^{(-i)}(X_i) + \frac{Z_i - \hat{e}^{(-i)}(X_i)}{\hat{e}^{(-i)}(X_i)(1 - \hat{e}^{(-i)}(X_i))} \left(Y_i - \hat{m}^{(-i)}(X_i) - \left(Z_i - \hat{e}^{(-i)}(X_i)\right)\hat{\tau}^{(-i)}(X_i)\right).
$$
\n(8)

BLP of CATE [\(Chernozhukov, 2018\)](https://arxiv.org/abs/1712.04802)

Conditional average treatment effect (CATE):

$$
g_0(1,X)-g_0(0,X)\\
$$

- captures (learnable) heterogeneity in treatment effects under unconfoundedness
- generally high-dimensional nonparametric object inference impractical (impossible?)

Another potential summary is best linear predictor (BLP) of CATE given pre-specified (low-dimensional) vector W

• other summaries possible; Semenova and Chernozhukov (2021)

Inference for BLP is possible using orthogonal score for ATE

6.1. Implementation Algorithm. We describe an algorithm based on the first identification strategy and provide some specific implementation details for the empirical example.

Algorithm 1 (Inference Algorithm). The inputs are given by the data on units $i \in [N] = \{1, ..., N\}$.

Step 0. Fix the number of splits S and the significance level α , e.g. $S = 100$ and $\alpha = 0.05$.

Step 1. Compute the propensity scores $p(Z_i)$ for $i \in [N]$.

Step 2. Consider S splits in half of the indices $i \in \{1, ..., N\}$ into the main sample, M, and the auxiliary sample, A. Over each split $s = 1, ..., S$, apply the following steps:

- a. Tune and train each ML method separately to learn $B(\cdot)$ and $S(\cdot)$ using A. For each $i \in M$, compute the predicted baseline effect $B(Z_i)$ and predicted treatment effect $S(Z_i)$. If there is zero variation in $B(Z_i)$ and $S(Z_i)$ add Gaussian noise with small variance to the proxies, e.g., a $1/20$ -th fraction of the sample variance of Y.
- b. Estimate the BLP parameters by weighted OLS in M , i.e.,

 $Y_i = \widehat{\alpha}' \underline{X_{1i}} + \widehat{\beta}_1 (D_i - p(Z_i)) + \widehat{\beta}_2 (D_i - p(Z_i))(S_i - \mathbb{E}_{N,M} S_i) + \widehat{\epsilon}_i, \ i \in M$

such that $\mathbb{E}_{N,M}[w(Z_i)\hat{\epsilon}_iX_i]=0$ for $X_i=[X'_{1i},D_i-p(Z_i),(D_i-p(Z_i))(S_i-\mathbb{E}_{N,M}S_i)]'$, where $w(Z_i) = \{p(Z_i)(1 - p(Z_i))\}^{-1}$ and X_{1i} includes a constant, $B(Z_i)$ and $S(Z_i)$.

c. Estimate the GATES parameters by weighted OLS in M , i.e.,

$$
Y_i = \widehat{\alpha}' X_{1i} + \sum_{k=1}^K \widehat{\gamma}_k \cdot (D_i - p(Z_i)) \cdot 1(S_i \in I_k) + \widehat{\nu}_i, \ i \in M,
$$

such that $\mathbb{E}_{N,M}[w(Z_i)\hat{\nu}_iW_i] = 0$ for $W_i = [X'_{i1}, \{(D_i - p(Z_i))1(S_i \in I_k)\}_{k=1}^K]$, where $w(Z_i) =$ $\{p(Z_i)(1-p(Z_i))\}^{-1}$, X_{1i} includes a constant, $B(Z_i)$ and $S(Z_i)$, $I_k = [\ell_{k-1}, \ell_k)$, and ℓ_k is the (k/K) -quantile of $\{S_i\}_{i \in M}$.

d. Estimate the CLAN parameters in M by

 $\widehat{\delta}_1 = \mathbb{E}_{N,M}[q(Y_i, Z_i) | S_i \in I_1]$ and $\widehat{\delta}_K = \mathbb{E}_{N,M}[q(Y_i, Z_i) | S_i \in I_K],$

where $I_k = [\ell_{k-1}, \ell_k)$ and ℓ_k is the (k/K) -quantile of $\{S_i\}_{i \in M}$.

Group Average Treatment Effects (GATEs)

Group average treatment effects (GATEs):

- Let G be an indicator for belonging to some group of interest (e.g. an education category)
- GATE = $E[g_0(1, X) g_0(0, X) | G = 1]$
- Can use to summarize heterogeneity along pre-specified directions of interest
- Average treatment effect on the treated (ATET) is a special case
- For $\psi_1(Y, Z, X)$ defined above, orthogonal moment for GATE is

$$
\mathrm{E}\left[\frac{G}{p_G}\psi_1(Y,Z,X)\right]=0
$$

• Nuisance functions: $E[Z|X] = m_0(X)$; $E[Y|Z,X] = g_0(Z,X)$; $p_G = E[G]$

- 1. Partition sample indices into random folds of approximately equal size: $\{1, ..., n\} = \bigcup_{k=1}^{K} I_k$. For each $k = 1, ..., K$, compute estimators $\hat{p}_{k1}, \hat{q}_{k1}$, and \hat{m}_{k1} of E[G] and the conditional expectation functions $g_0(Z, X) = E[Y|Z, X]$ and $m_0(X) = E[Z|X]$ leaving out the k^{th} block of data and enforcing $\epsilon \leq \widehat{m}_{[k]} \leq 1 - \epsilon$.
- 2. For each $i \in I_k$, let

$$
\widehat{\psi}(Y_i, Z_i, X_i, G_i; \alpha) = \frac{G_i}{\widehat{\rho}_{[k]}} \left(\widehat{g}_{[k]}(1, X_i) - \widehat{g}_{[k]}(0, X_i) + \frac{Z_i(Y_i - \widehat{g}_{[k]}(1, X_i))}{\widehat{m}_{[k]}(X_i)} - \frac{(1 - Z_i)(Y_i - \widehat{g}_{[k]}(0, X_i))}{1 - \widehat{m}_{[k]}(X_i)} \right) - \frac{G_i}{\widehat{\rho}_{[k]}} \alpha.
$$

Compute the estimator $\widehat{\alpha}$ as the solution to $\mathbb{E}_n[\widehat{\psi}(W_i;\alpha)] = 0$ which yields

$$
\widehat{\alpha} = \frac{\mathbb{E}_n \left[\frac{G_i}{\widehat{P}_{[k]}} \left(\widehat{g}_{[k]}(1, X_i) - \widehat{g}_{[k]}(0, X_i) + \frac{Z_i(Y_i - \widehat{g}_{[k]}(1, X_i))}{\widehat{m}_{[k]}(X_i)} - \frac{(1 - Z_i)(Y_i - \widehat{g}_{[k]}(0, X_i))}{1 - \widehat{m}_{[k]}(X_i)} \right) \right]}{\mathbb{E}_n \left[\frac{G_i}{\widehat{P}_{[k]}} \right]}.
$$

3. Let

$$
\widehat{\varphi}(Y_i, Z_i, X_i, G_i) = \frac{\widehat{\psi}(Y_i, Z_i, X_i, G_i; \widehat{\alpha})}{\mathbb{E}_n \left[\frac{G_i}{\widehat{P}_{[k]}}\right]}.
$$

Construct standard errors via

$$
\sqrt{\hat{V}}/n, \quad \hat{V} = \mathbb{E}_n[\hat{\varphi}(Y_i, Z_i, X_i, G_i)^2]
$$

and use standard normal critical values for inference.

- Variable importance
	- CF-noncluster: [1] "pretest" "Q34_1_feb" "Q13_8_feb_2" "Q25_apr" "EOC_scale_score"
	- Cluster-robust CF: [1] "pretest" "EOC_scale_score" "absent_days" "EOC_achieve_level" "mean_num_received"

Permutation Importance

